Hypothesis testing with e-values

ISI WSC IPS Discussion

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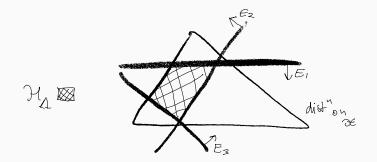
- Practical, intuitive toolbox for designing things that work.
- Beautiful open questions in the theory.

GRO Non-parametrics

Larsson constructs hypotheses by requiring a family $(E_{\lambda})_{\lambda \in \Lambda}$ to be e-values:

e-value hypothesis class

$$\mathcal{H}_{\Lambda} := \left\{ P \text{ on } \mathcal{X} \mid \forall \lambda \in \Lambda : \mathbb{E}_{X \sim P} \left[E_{\lambda}(X) \right] \leq 1 \right\}$$



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Example

Interesting example: bounded $(1 + \epsilon)$ -th moment for $\epsilon > 0$:

$$\mathbb{E}_{X \sim P}\left[|X|^{1+\epsilon}\right] \leq B$$

Heavy-tailed distributions including Pareto, Fisher-Tippett, ...

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GRO $S^* := \underset{S \text{ an e-value for } \mathcal{H}_0}{\operatorname{arg max}} \mathbb{E}_{X \sim Q} \left[\ln S(X) \right]$

Question

What does the non-parametric GRO look like?

Larsson's theorem (finite Λ version)

Theorem

S is an e-value for \mathcal{H}_{Λ} if there are non-negative $\pi \geq 0$ such that

$$S(x) \leq 1 + \sum_{\lambda \in \Lambda} \pi_{\lambda}(E_{\lambda}(x) - 1)$$
 for all $x \in \mathcal{X}$.

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The GRO e-value for \mathcal{H}_Λ is found by solving

$$\max_{S \text{ an e-value for } \mathcal{H}_0} \mathbb{E}_{X \sim Q} \left[\ln S(X) \right] = \max_{\pi \geq 0} \mathbb{E}_{X \sim Q} \left[\ln \left(1 + \sum_{\lambda \in \Lambda} \pi_\lambda \left(E_\lambda(X) - 1 \right) \right) \right]$$

We may also write the following

$$\min_{\substack{P \in \mathcal{H}_{\Lambda} \\ \pi \ge 0}} \operatorname{KL}(Q \| P) = \max_{\substack{\nu \in \mathbb{R} \\ \pi \ge 0}} \min_{P \ge 0} \mathbb{E}_{Q} \left[\ln \frac{Q(X)}{P(X)} \right] + \sum_{\lambda \in \Lambda} \pi_{\lambda} \left(\mathbb{E}_{X \sim P} \left[E_{\lambda}(X) - 1 \right] \right) + \nu \left(\mathbb{E}_{X \sim P} [1] - 1 \right)$$

So that

$$P(x) = rac{Q(x)}{
u + \sum_{\lambda \in \Lambda} \pi_{\lambda} E_{\lambda}(x)}$$

All in all, the GRO e-value is a likelihood ratio

$$S^*(x) = rac{Q(x)}{P_{
u^*,\lambda^*}(x)}$$

And we are back in the **Turner 2021** case.

Techniques partially developed/exploited in Bandit literature under the names **KLInf** and **empirical likelihood**

(Honda and Takemura, 2010; Cappé et al., 2013; Agrawal, Juneja, and Glynn, 2020; Agrawal, Juneja, and Koolen, 2021; Agrawal,

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Question

Non-linear conditions? Variance? CVaR? Centered moment constraints?

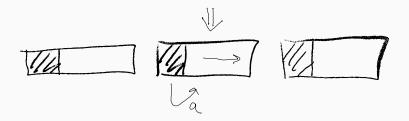
Choosing the Alternative

Vovk

Cox' problem: test all means 0 vs one mean $\mu \neq 0$.

Vovk proposes

- Use mean on first blocks to pick a population.
- Compute the mean (ML) *a* of the first block.
- Use i.i.d. *P*_a as the alternative model for the second block.



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Consideration:

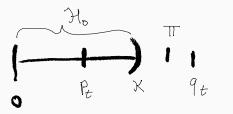
- The first blocks are used to select a block: *a* will overfit (slightly).
- If $a > \mu$ then $E = P_a/P_0$ is only expected to win if $a < 2\mu$.
- Are there ways to dampen a?
- Curiously: problem gets **worse** if there are *multiple* populations with $\mu \neq 0$.

Henzi

At each time-point t, the hypothesis is that p_t is a better prediction that q_t (as measured by score function S).

Henzi proposes (say $p_t < q_t$)

- Compute the mid-point $\kappa \in (p_t, q_t)$ (halfway for Brier).
- Use null $\mathcal{H}_0 = \{ \mathsf{Ber}(\theta) \mid \theta \in [0, \kappa] \}.$
- Use alternative $\pi = \frac{3}{4}q_t + \frac{1}{4}p_t > \kappa$.
- Multiply evidence by e-value $P_{\pi}(Y)/P_{\kappa}(Y)$.



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Consideration:

- π is expected to gain evidence for true parameter μ between κ and π. But what if μ = κ + ε?
- Are there (effect size?) guidelines for setting ϵ a priori?
- Universal modelling (mixture) over ϵ ?

Safety and Power

Joy and elegance

Sequential story only partially understood.

- Product of GRO may not be GRO itself. *Order of quantifiers matters!*
- GRO+invariance sometimes leads to test-super-martingale in reduced filtration.
 When? Approximately?

Exciting area!

Thanks!

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